

Influence Analysis with Panel Data using STATA

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Motivation

- ▶ Short panel data sets (small N but $N \gg T$) are common in many fields of Economics
 - ▶ Macro-level panel data (e.g., 50 US States)
 - ▶ Cell-group data (e.g., gender-age-occupation)
 - ▶ Experimental panel data (e.g., limited no. participants)
- ▶ Observational data may contain “anomalous” observations (Rousseeuw and Van Zomeren, 1990; Silva, 2001)
 - ▶ Vertical outliers (VO), good leverage (GL) points, bad leverage (BL) points ▶ Example ▶ DGP
- ▶ Large influence on the Least Squares (LS) estimates
⇒ Biased regression coefficients or standard errors (Donald and Maddala, 1993; Bramati and Croux, 2007; Verardi and Croux, 2009)

Motivation

- ▶ **Diagnostic plots** (leverage-vs-residual plots)
 - ▶ for cross-sectional data: `lvr2plot/lvr2plot2`
 - ▶ Less handy for panel data
- ▶ **Measures of influence** (Cook (1979)'s distance)
 - ▶ for cross-sectional data:
`predict c, cooks`
 - ▶ for panel data:
`jackknife2, cooks(newvar) bpd(newvar): command`
 - ▶ These metrics may fail to flag multiple atypical cases (Atkinson and Mulira, 1993; Chatterjee and Hadi, 1988; Rousseeuw and Van Zomeren, 1990) unlike *pair-wise measures* (Lawrance, 1995)

In this talk

- ▶ I present a method to
 1. **Detect** anomalous units and **identify** their type
 2. **Show** how these affect the LS estimates
- ▶ I follow a *unit-wise* approach (full history of a unit)
- ▶ I develop two commands in Stata
 - ▶ `xtlvr2plot` – Leverage-vs-residual plot for panel data
 - ▶ `xtinfluence` – Pair-wise influence measures with panel data
- ▶ I apply the method to a cross-country study
 - ▶ Berka et al. (2018, AER)

Model and estimator

Static linear panel regression model with fixed effects

$$y_{it} = \mathbf{x}_{it}'\boldsymbol{\beta} + \alpha_i + u_{it}$$

Model after the *within-group* (WG) transformation

$$\tilde{y}_{it} = \tilde{\mathbf{x}}_{it}'\boldsymbol{\beta} + \tilde{u}_{it}$$

where $\tilde{y}_{it} = y_{it} - T^{-1} \sum_t y_{it}$, etc., and $\boldsymbol{\beta}$ is a vector of parameters.

The WG Estimator

$$\hat{\boldsymbol{\beta}} = \left(\sum_{i=1}^N \sum_{t=1}^T \tilde{\mathbf{x}}_{it} \tilde{\mathbf{x}}_{it}' \right)^{-1} \sum_{i=1}^N \sum_{t=1}^T \tilde{\mathbf{x}}_{it} \tilde{y}_{it}$$

Overview

1. Identify anomalous units with `xtlvr2plot`
2. Understand how anomalous units may affect the LS estimates with `xtinfluence`
 - 2.1 Joint influence and joint effect
 - 2.2 Conditional influence and conditional effect

xtlvr2plot: Syntax

xtlvr2plot – Leverage-versus-normalised residual squared plot for panel data.
xtset 'panelvar' 'timevar' is required.

```
xtlvr2plot depvar [indepvar] [if] [in] [, options]
```

options

graph_opts

graph options allowed for twoway scatter

Generated variables

lev

average individual leverage

normres2

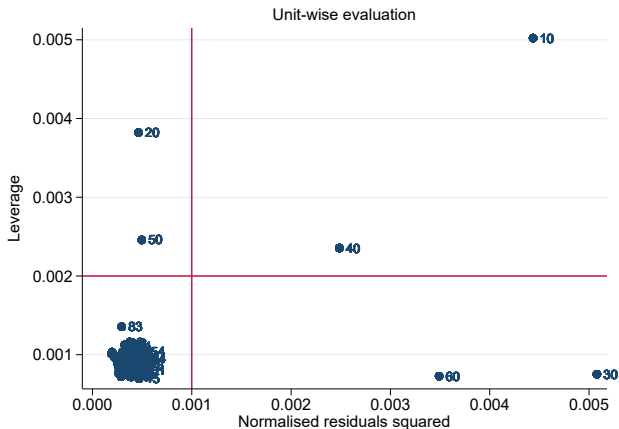
average individual residual squared

xtlvr2plot: Example

```
** Use of the 'xtlvr2plot' command
xtset id time

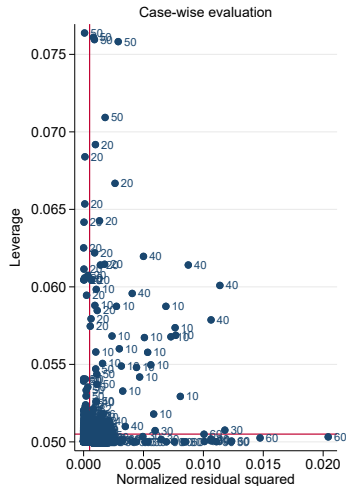
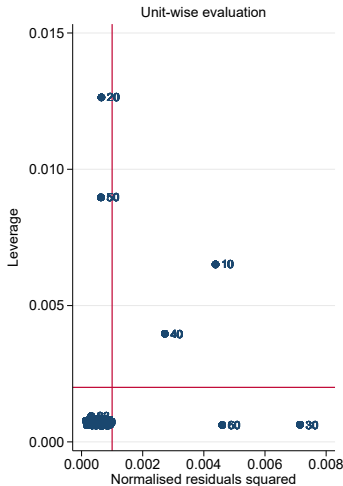
xtlvr2plot y x,          ///
    mlabel(id)          ///
    xlabel(, format(%9.3fc))  ///
    ylabel(, angle(h) format(%9.3fc))  ///
    title("Unit-wise Evaluation", size(medsmall))  ///
    saving("xtlvr2plot_example.gph", replace)
```

xtlvr2plot: Leverage-vs-residual plot



Note: Units 10 and 40 are set to be bad leverage units; units 20 and 50 good leverage units; units 30 and 60 vertical outliers. [► Formulae](#)

xtlvr2plot vs lvr2plot



Note: Units 10 and 40 are set to be bad leverage units; units 20 and 50 good leverage units; units 30 and 60 vertical outliers.

xtlvr2plot: Summary Table

```
** Summary table w/detected anomalous units  
** generated by 'xtlvr2plot'
```

Anomalous units	
x-cutoff =	0.001
y-cutoff =	0.002
Good leverage units	
- Count :	2
- List :	20 50
Bad leverage units	
- Count :	2
- List :	10 40
Vertical outliers	
- Count :	2
- List :	30 60

Note: Units 10 and 40 are set to be bad leverage units; units 20 and 50 good leverage units; units 30 and 60 vertical outliers.

Overview

1. Identify anomalous units with `xtlvr2plot`
2. Understand how anomalous units may affect the LS estimates with `xtinfluence`
 - 2.1 Joint influence and joint effect
 - 2.2 Conditional influence and conditional effect

Joint measures

- ▶ For $i \neq j$, joint influence is

$$C_{ij}(\hat{\beta}) = (\hat{\beta} - \hat{\beta}_{(i,j)})' (\tilde{\mathbf{X}}' \tilde{\mathbf{X}}) (\hat{\beta} - \hat{\beta}_{(i,j)}) (s^2 K)^{-1}$$

where $\hat{\beta}_{(i,j)}$ is WG estimator w/t units i and j , s is RMSE, K is no. covariates

- ▶ Influence exerted by a pair (i,j) on LS estimates *jointly*
- ▶ Comparison of LS estimates *with* and *without* the *pair*
- ▶ For $i = j$, i 's individual influence (as in [Belotti and Peracchi \(2020\)](#))

- ▶ The joint effect is

$$K_{j|i} = \frac{C_{ij}(\hat{\beta})}{C_{ii}(\hat{\beta})}$$

- ▶ How much the pair is influential wrt i
- ▶ For large values of $K_{j|i}$, j *alters* the effect of i
 - ▶ j either *enhances* or *reduces* the effect of i on the LS estimates, based on the conditional effect

Joint measures

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 - ▶ j either *enhances* or *reduces* the effect of i on the LS estimates, based on the conditional effect

Conditional measures

- ▶ For $i \neq j$, conditional influence is

$$C_{i(j)}(\hat{\beta}) = (\hat{\beta}_{(i,j)} - \hat{\beta}_{(j)})' \left(\sum_{\substack{i=1 \\ i \neq j}}^N \tilde{\mathbf{X}}_{i(j)}' \tilde{\mathbf{X}}_{i(j)} \right) (\hat{\beta}_{(i,j)} - \hat{\beta}_{(j)}) (s^2 K)^{-1}$$

- ▶ Influence exerted by i on LS estimates without j in the sample
- ▶ How the absence of j affects the influence i on LS estimates
- ▶ The conditional effect is

$$M_{i(j)} = \frac{C_{i(j)}(\hat{\beta})}{C_{ii}(\hat{\beta})}$$

- ▶ How influence of i changes before and after the deletion of j
- ▶ If $M_{i(j)} \geq 1$, influence of i **increases** without j in the sample
 $\implies j$ **masks** i .
- ▶ If $M_{i(j)} < 1$, influence of i **decreases** without j in the sample
 $\implies j$ **boosts** i

Conditional measures

- ▶ For $i \neq j$, conditional influence is

$$C_{i(j)}(\hat{\beta}) = (\hat{\beta}_{(i,j)} - \hat{\beta}_{(j)})' \left(\sum_{\substack{i=1 \\ i \neq j}}^N \tilde{\mathbf{X}}_{i(j)}' \tilde{\mathbf{X}}_{i(j)} \right) (\hat{\beta}_{(i,j)} - \hat{\beta}_{(j)}) (s^2 K)^{-1}$$

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xtinfluence: Syntax

`xtinfluence` – Calculates and displays joint and conditional measures/effects of pairs of units i and j . The size of the symbols is proportional to the magnitude of the calculated measures. `xtset` ‘`panelvar`’ ‘`timevar`’ is required.

`xtinfluence` *depuvar* [*indepvar*] [*if*] [*in*] [, *options*]

options

figure(*graphtype*)

displays diagnostic plots as *graphtype*. Allowed *graphtype* are scatter plot or heat plot; default is scatter

graph_opts

saving(*filename*)

graph options allowed for scatter and heatmap
saves .dta and .pdf file with the specified name and location

Generated variables

_newid

assigns a new numeric identifier to sorted ‘`panelvar`’

Saved data sets

filename_adj_mtx.dta

Automatically saves a data set with the influence measures and effects generated by the command

xtinfluence: Example

****Use of the 'xtinfluence' command**

```
xtset id t
```

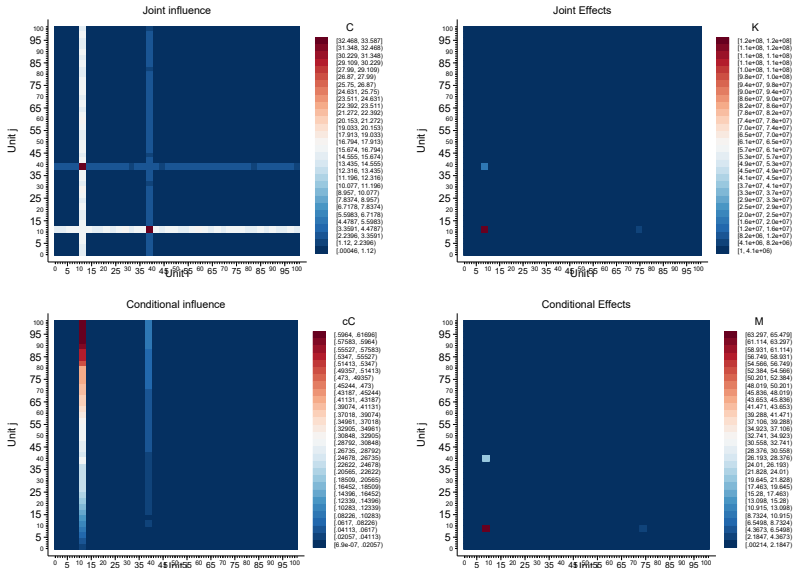
**** Heat plot**

```
xtinfluence y x, figure(heat)           ///  
    keylabels(all, interval) color(RdBu, reverse)  ///  
    lev(30) statistic(max)                ///  
    xlabel(5(10)100, angle(h) labsize(small))    ///  
    xmtick(##10) xlabel(##2, angle(h))           ///  
    ylabel(5(10)100, angle(h))                  ///  
    ymtick(##10) ylabel(##2, angle(h))           ///  
    saving("xtinfluence_heat")
```

**** Scatter plot**

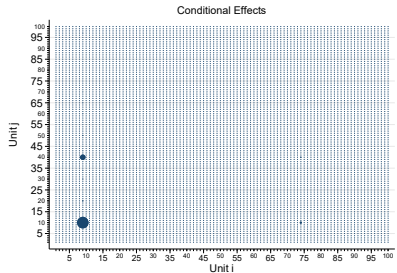
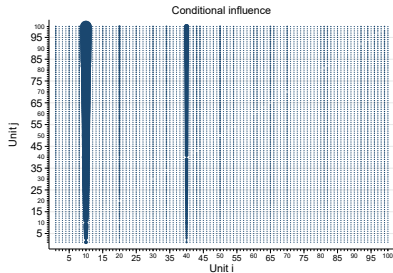
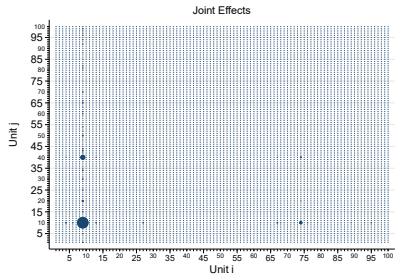
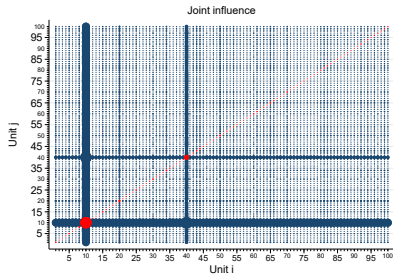
```
xtinfluence y x, figure(scatter)         ///  
    xlabel(5(10)100, angle(h) labsize(small))    ///  
    xmtick(##10) xlabel(##2, angle(h))           ///  
    ylabel(5(10)100, angle(h))                  ///  
    ymtick(##10) ylabel(##2, angle(h))           ///  
    saving("xtinfluence_scatter")
```

xtinfluence: Plot



Note: Units 10 and 40 are bad leverage units; units 20 and 50 good leverage units; units 30 and 60 vertical outliers. [▶ Adj-mtx](#)

xtinfluence: Plot



Note: Units 10 and 40 are bad leverage units; units 20 and 50 good leverage units; units 30 and 60 vertical outliers. [▶ Adj-mtx](#)

xtinfluence: Summary table

Variable	Obs	Mean	Std. dev.	Min	Max
C	10,000	.3811386	2.200585	2.35e-11	33.58732
K	10,000	16156.08	1242556	4.42e-08	1.23e+08
cC	10,000	.0038312	.0353837	0	.6169614
M	9,900	.0305928	.6922132	4.39e-06	65.47916

Influence analysis

v1 = k+1 = 2

v2 = NT-N-k-1 = 1898

c1 = 4/N = .04

c2 = F(v1,v2,.5) = 0.6934

Cii >= c1

- Count : 8

- List : 8 10 20 34 40 43 50 65

Cii >= c2

- Count : 2

- List : 10 40

i with K >= p99

- Count : 30

- List : 3 4 6 9 11 13 14 19 24 27 47 49 55 57 62 64 67 68 69 71 72 74 76 77 79 84 86 89 93 95

j with K >= p99

- Count :

- List :

i with M >= 1

- Count : 2

- List : 9 74

j with M >= 1

- Count : 2

- List : 10 40

Empirical example

- ▶ [Berka et al. \(2018, AER\)](#) – Macro-panel data
 - ▶ Objective: Study relationship between real exchange rate and sectoral productivity in the Eurozone
 - ▶ Units of observation: 9 EU countries
 - ▶ Time period: 1995–2007

▶ Overview

▶ Scatter

▶ `xtlvr2plot`

▶ `xtinfluence`

▶ Summary

Conclusion

- ▶ The proposed STATA commands allow to systematically
 1. Identify anomalous units and their type (unit-wise leverage-vs-residual plot)
 2. Investigate how anomalous units may affect the LS estimates (joint and conditional influence and effects)
- ▶ Once the type of anomaly is identified, the literature suggests, e.g.,
 1. Methods for measurement error if error in the data entry
 2. Robust estimation techniques if VO and BL ([Bramati and Croux, 2007](#); [Verardi and Croux, 2009](#); [Aquaro and Čížek, 2013, 2014](#); [Jiao, 2022](#))
 3. Jackknife-type standard errors if GL ([MacKinnon and White, 1985](#); [Davidson et al., 1993](#); [MacKinnon, 2013](#); [Belotti and Peracchi, 2020](#); [Polselli, 2022](#))

Thank you for your attention!

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🐙 <https://github.com/POLSEAN/Influence-Analysis>

🐦 [@AnnalivPolselli](https://twitter.com/AnnalivPolselli)

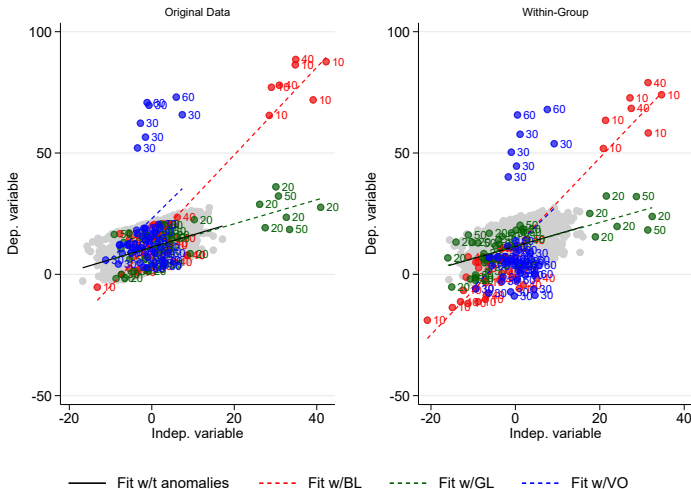
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References II

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- MacKinnon, J. G. and White, H. (1985). Some heteroskedasticity-consistent covariance matrix estimators with improved finite sample properties. *Journal of econometrics*, 29(3):305–325.
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Scatter Plot DGP

[▶ Back](#)

Note: Units 10 and 40 are bad leverage units; units 20 and 50 are good leverage units; units 30 and 60 are vertical outliers.

```
loc numobs 100
set obs 100
gen id = _n
expand 20

bys id: generate t = _n
bys id: gen z = rnormal(0,5)
**GL
bys id: replace z = z + rnormal(30,1) if id==20 & t<=5
bys id: replace z = z + rnormal(30,1) if id==50 & t<=2
**for BL
bys id: replace z = z + rnormal(30,1) if id==10 & t<=5
bys id: replace z = z + rnormal(30,1) if id==40 & t<=2
**line
bys id: gen a = runiform(0,20)
bys id: gen y = 1 + .5*z + a + runiform()
**BL
bys id: replace y = y + rnormal(50,1) if id==10 & t<=5
bys id: replace y = y + rnormal(50,1) if id==40 & t<=2
*V0
bys id: replace y = y + rnormal(50,1) if id==30 & t<=5
bys id: replace y = y + rnormal(50,1) if id==60 & t<=2
```

Example: Berka et al. (2018) [▶ Back](#)

- ▶ They study relationship between real exchange rate and sectoral productivity in the Eurozone
- ▶ Regression model:

$$RER_{it} = \beta TFP_{it} + \mathbf{x}_{it}'\boldsymbol{\gamma} + \alpha_i + u_{it}$$

RER_{it} : real exchange rate in log

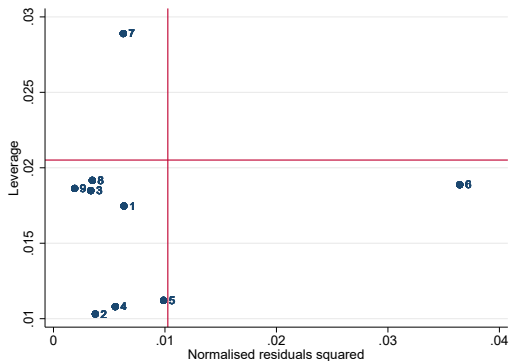
TFP_{it} : total factor productivity in log

$\mathbf{x}_{it}$: other controls

α_i : country fixed effects

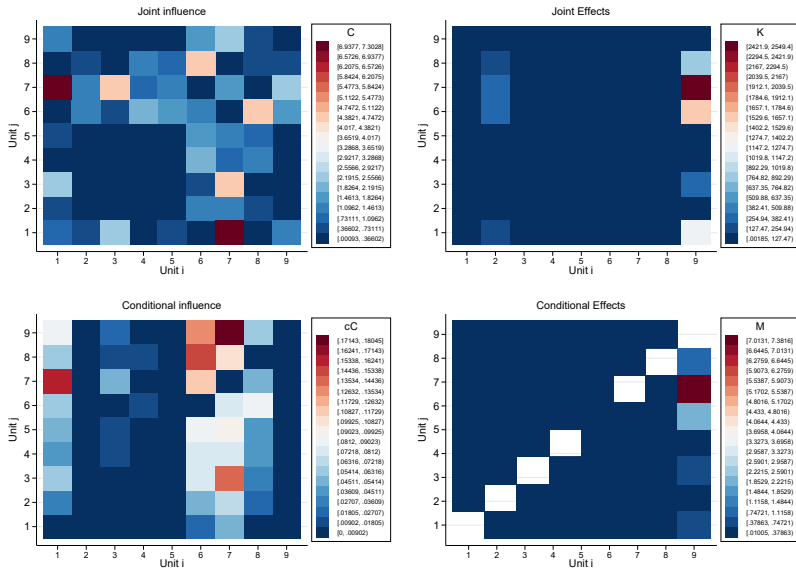
- ▶ Finding strong correlation between TFP and RER among high-income countries with floating nominal exchange rates
- ▶ Sample: 9 EU countries
- ▶ Time Period: 1995–2007
- ▶ Table 4, specification (2a)

Example: Leverage-vs-residual plot [▶ Back](#)



Note: 1-Austria, 2-Belgium, 3-Finland, 4-France, 5-Germany, 6-Ireland, 7-Italy, 8-Netherlands, 9-Spain.

Example: Network-like plots [▶ Back](#)



Example: Summary [▶ Back](#)

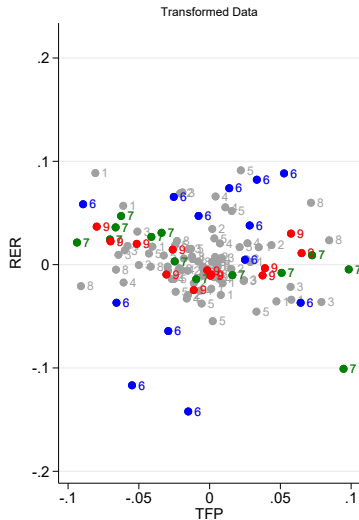
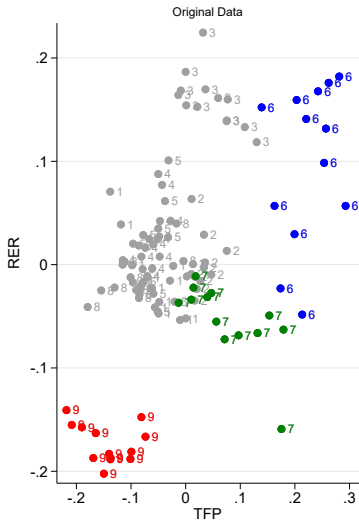
Variable	Obs	Mean	Std. dev.	Min	Max
C	81	1.0233	1.472976	.0009253	7.30281
K	81	97.87085	368.2484	.0018538	2549.404
cC	81	.032125	.0439157	0	.1804506
M	72	.2303033	.8915019	.0046645	7.381636

Influence analysis

```
v1 = k+1 = 2
v2 = NT-N-k-1 = 184
c1 = 4/N = .4444444444444444
c2 = F(v1,v2,.5) = 0.6958
```

```
Cii >= c1
- Count : 4
- List : 1 6 7 8
Cii >= c2
- Count : 3
- List : 1 6 7
i with K >= p99
- Count : 1
- List : 9
j with K >= p99
- Count :
- List :
i with M >= 1
- Count : 1
- List : 9
j with M >= 1
- Count : 2
- List : 6 7
```

Example: Scatter [▶ Back](#)



Note: 1-Austria, 2-Belgium, 3-Finland, 4-France, 5-Germany, 6-Ireland, 7-Italy, 8-Netherlands, 9-Spain.

filename_adj_mtx.dta [▶ Back](#)

The saved data set resembles a directed and weighted adjacency list

	i	j	C	K	cC	M
1	1	1	.0318985	1	0	0
2	1	2	.0779802	2.444638	8.05e-06	.0002523
3	1	3	.0379366	1.189292	.000065	.0020391
4	1	4	.0012006	2.545595	.0000804	.0025191
5	1	5	.0384888	1.206603	.0000916	.0028703
6	1	6	.0619195	1.941144	.000091	.0028528
7	1	7	.0802803	2.516744	.0001116	.0034988
8	1	8	.0322271	1.010302	.0001236	.003874
9	1	9	.0102966	.3227937	.0001144	.0035852
10	1	10	34.86443	1092.981	.0001167	.0036569
11	1	11	.0380862	1.193983	.0001264	.0039615
12	1	12	.0524164	1.643225	.0001519	.0047621
13	1	13	.0510088	1.599099	.0001667	.005226
14	1	14	.0550416	1.725525	.0001834	.0057488
15	1	15	.0617752	1.936618	.0001679	.0052648
16	1	16	.0591808	1.855285	.000202	.0063336
17	1	17	.0512263	1.605917	.0001969	.0061739
18	1	18	.067513	2.116496	.0002049	.006424
19	1	19	.0904264	2.834818	.000237	.0074296
20	1	20	11.59427	363.474	.0005592	.0175295
21	1	21	.0564583	1.769938	.0002562	.0080332
22	1	22	.0020566	.0644732	.0002375	.0074454
23	1	23	.091529	2.869384	.0002585	.0081049
24	1	24	.026083	.8176892	.0002669	.0083674
25	1	25	.0945991	2.965631	.0003046	.0095503

Residual and Leverage

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- ▶ The **average normalised residual** squared

$$\hat{u}_i^* = \frac{1}{T} \sum_{t=1}^T \left(\frac{\hat{u}_{it}}{\sqrt{\sum_i \hat{u}_{it}^2}} \right)^2$$

where $\hat{u}_{it} = \tilde{y}_{it} - \tilde{\mathbf{x}}'_{it} \hat{\boldsymbol{\beta}}$ are LS Residuals.

Cut-off value: $c_{\hat{u}_i^*} = \frac{2}{NT}$

- ▶ The **average individual leverage** of unit i at time t is

$$\bar{h}_i = \frac{1}{T} \sum_{t=1}^T h_{ii,tt}$$

where $h_{ii,tt} = \tilde{\mathbf{x}}'_{it} (\tilde{\mathbf{X}}' \tilde{\mathbf{X}})^{-1} \tilde{\mathbf{x}}_{it}$, and $h_{ii,ts} = \tilde{\mathbf{x}}'_{it} (\tilde{\mathbf{X}}' \tilde{\mathbf{X}})^{-1} \tilde{\mathbf{x}}_{is}$ for $t, s = 1, \dots, T$.

Cut-off value: $c_{\bar{h}_i} = \frac{2(K+1)}{NT}$

Summary of method

1. Identify anomalous units and their type with `xtlvr2plot`
2. Conduct the influence analysis with `xtinfluence`
 - 2.1 **Joint Influence Plot**
 - Identify units with high individual influence (main diagonal)
 - Identify pairs with high joint influence (off-diagonal)
 - Highly influential units swamp all other units
 - 2.2 **Joint Effect Plot**
 - Identify pairs with largest effect
 - j swamps the effect of i
 - j must be detected in (1) and (2.1)
 - 2.3 **Conditional Influence Plot**
 - Identify influential i conditional to removing j
 - Check if same units as (1) and (2.1)
 - 2.4 **Conditional Effect Plot**
 - Identify pairs with largest effect
 - j masks the effect of i
 - Compare identified pairs with (2.2)
3. Units detected in (1), (2.1) and (2.3) are anomalous; (2.2) and (2.4) explain how they affect the influence of other units and, hence, LS estimates